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# Ionising radiation and cardiovascular disease: systematic review and meta-analysis

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## ABSTRACT

### OBJECTIVE

To systematically review and perform a meta-analysis of radiation associated risks of cardiovascular disease in all groups exposed to radiation with individual radiation dose estimates.

### DESIGN

Systematic review and meta-analysis.

### MAIN OUTCOME MEASURES

Excess relative risk per unit dose (Gy), estimated by restricted maximum likelihood methods.

### DATA SOURCES

PubMed and Medline, Embase, Scopus, Web of Science Core collection databases.

### ELIGIBILITY CRITERIA FOR SELECTING STUDIES

Databases were searched on 6 October 2022, with no limits on date of publication or language. Animal studies and studies without an abstract were excluded.

### RESULTS

The meta-analysis yielded 93 relevant studies. Relative risk per Gy increased for all cardiovascular disease (excess relative risk per Gy of 0.11 (95% confidence interval 0.08 to 0.14)) and for the four major subtypes of cardiovascular disease (ischaemic heart disease, other heart disease, cerebrovascular disease, all other cardiovascular disease). However, interstudy heterogeneity was noted ( $P < 0.05$  for all

endpoints except for other heart disease), possibly resulting from interstudy variation in unmeasured confounders or effect modifiers, which is markedly reduced if attention is restricted to higher quality studies or those at moderate doses ( $< 0.5$  Gy) or low dose rates ( $< 5$  mGy/h). For ischaemic heart disease and all cardiovascular disease, risks were larger per unit dose for lower dose (inverse dose effect) and for fractionated exposures (inverse dose fractionation effect). Population based excess absolute risks are estimated for a number of national populations (Canada, England and Wales, France, Germany, Japan, USA) and range from 2.33% per Gy (95% confidence interval 1.69% to 2.98%) for England and Wales to 3.66% per Gy (2.65% to 4.68%) for Germany, largely reflecting the underlying rates of cardiovascular disease mortality in these populations. Estimated risk of mortality from cardiovascular disease are generally dominated by cerebrovascular disease (around 0.94-1.26% per Gy), with the next largest contribution from ischaemic heart disease (around 0.30-1.20% per Gy).

### CONCLUSIONS

Results provide evidence supporting a causal association between radiation exposure and cardiovascular disease at high dose, and to a lesser extent at low dose, with some indications of differences in risk between acute and chronic exposures, which require further investigation. The observed heterogeneity complicates a causal interpretation of these findings, although this heterogeneity is much reduced if only higher quality studies or those at moderate doses or low dose rates are considered. Studies are needed to assess in more detail modifications of radiation effect by lifestyle and medical risk factors.

### SYSTEMATIC REVIEW REGISTRATION

PROSPERO CRD42020202036

### Introduction

Cardiovascular diseases are the leading cause of death worldwide.<sup>1 2</sup> Cardiovascular disease was the underlying cause of death for about a third of the 2.8 million deaths in the USA in 2018: ischaemic heart disease accounted for 42% and stroke for 17% of all cardiovascular disease deaths.<sup>3</sup> Worldwide, ischaemic heart disease ranks first in years of life lost and stroke ranks third. Consistently identified independent risk factors include age, smoking, diabetes mellitus, hypertension, obesity, and increased total and low density lipoprotein or decreased high density lipoprotein cholesterol.<sup>4-6</sup> A heritable genetic

## WHAT IS ALREADY KNOWN ON THIS TOPIC

Exposure to high dose ionising radiation during radiotherapy can damage the heart

Cardiovascular disease risk in the low dose range ( $< 0.1$  Gy), characteristic of doses that patients receive from medical diagnostic exposures or those radiation workers receive from occupational exposures is not well understood

Previous systematic reviews published over a decade ago looked at a much smaller number of studies, mostly with lower dose or lower dose rate exposures

## WHAT THIS STUDY ADDS

A systematic review of 15 098 studies yielded 93 informative and largely non-overlapping studies and suggest modest but significantly increased excess lifetime risk of 2.3-3.9 deaths per 100 people exposed to one Gy of radiation

These findings have implications for patients who undergo radiation exposure as part of their medical care, as well as policy makers involved in managing radiation risks to radiation workers and the public

The potential increased risk of radiogenic cardiovascular disease should prompt vigilance to control other modifiable cardiovascular risk factors and extra consideration of cardiovascular disease following radiation exposure

component for coronary heart disease has also been reported.<sup>7-10</sup>

Environmental factors might also contribute to cardiovascular disease risk and exposure to ionising radiation during radiotherapy can damage the heart.<sup>11</sup> Radiotherapy doses to the heart and other organs or tissues of relevance to the cardiovascular system can be very high, with doses to some regions of the heart exceeding 40 Gy in previous years;<sup>12</sup> although doses tend to be lower among groups treated for non-malignant disease than for cancer, and lower among people treated for cancer in more recent years.<sup>13</sup> Many older studies of radiotherapy and cardiovascular disease do not have detailed individual radiation organ dosimetry,<sup>14-18</sup> or data for concomitant chemotherapy drugs, of which some types (eg, vinca alkaloids including vincristine, and anthracyclines including doxorubicin) are cardiotoxic, irrespective of the administration of concomitant radiotherapy.<sup>17</sup> Concomitant chemotherapy is often correlated with radiotherapy dose therefore confounding of the dose response is possible.

The Life Span Study of the Japanese atomic bomb survivors provides evidence of increased risk of cardiovascular disease at lower levels of dose, less than 5 Gy, and with mean doses of much less than 0.5 Gy.<sup>19-20</sup> No findings suggested an appreciable non-linear association in the radiation dose-response for cardiovascular disease mortality in the Life Span Study data, although the form of the dose-response relation, particularly at doses less than 0.5 Gy, is uncertain.<sup>20</sup> Therefore, the extent of cardiovascular disease risk is uncertain for low doses (<0.1 Gy), which are characteristic of doses from medical diagnostic exposures. Emerging, and still controversial, evidence suggests that exposure to much lower doses and dose rates of radiation, in particular occupational and medical diagnostic exposure,<sup>21</sup> might be associated with excess risk of cardiovascular disease. Claims have been made of a no effect dose threshold for cardiovascular disease mortality in the Life Span Study, below which no radiation induced excess risk exists,<sup>22</sup> although this finding has been disputed.<sup>23</sup> Observational epidemiological studies are likely to have difficulty in detecting increased risk at low dose levels because the main types of cardiovascular disease of concern are very common in the population as a whole and because of the multiple contributory risk factors that are potentially confounding. The International Commission on Radiological Protection has classified cardiovascular disease as a tissue reaction (formerly termed a deterministic effect), with an approximate nominal threshold dose of 0.5 Gy independent of dose rate.<sup>24</sup> This level is determined by linear models fitted to epidemiological data that yield less than a 1% lifetime risk. As such, this threshold is a practical one but is not a true no effect dose threshold.

In this systematic review and meta-analysis, we research the risks of radiation associated cardiovascular disease that have been observed in therapeutically or diagnostically exposed cohorts. Risks among

groups exposed to generally lower levels of radiation dose (with maximum dose <0.5 Gy) or low dose rate (<5 mGy/h) are also assessed, specifically in the Life Span Study and in groups that are occupationally and environmentally exposed. Attention is concentrated on studies with informative individual organ dosimetry. In contrast to previous systematic reviews,<sup>21-25-26</sup> which were published at least 10 years ago, we do not limit our inclusion to the lower dose literature; a previous review and meta-analysis covered literature up to about 2016, but the review was not systematic.<sup>27</sup>

## Methods

### Selection of studies

We conducted a systematic review and reported according to PRISMA and registered in PROSPERO (CRD42020202036) on 1 November 2020. Thereafter, we made a few small changes in the course of progress with the screening, which are detailed in PROSPERO. We used PubMed/Medline, Embase, Scopus, and Web of Science's Core Collection to systematically search the literature, with no limits applied (date, language), on 6 October 2022. Cardiovascular disease is defined as those causes of mortality and incidence with International Classification of Diseases 10th revision (ICD-10) cardiovascular disease codes I00-I99 (or equivalently the ICD 8th revision codes 390-458 or ICD 9th revision (ICD-9) codes 390-459). We excluded animal studies and any study without an abstract. The database search was conducted by AL with input from MPL and NH, and yielded a total of 15 098 articles; these were loaded into Covidence by AL and then subjected to joint review by MPL and NH. In the first stage, we used only title and abstract to determine eligibility. Later stages of the search used much more complete information. Among other things, the reviewers (MPL, NH) independently ascertained if the organ dosimetry was adequate to estimate radiation risk and to determine if they were potentially informative on the desired outcome measure, ie, excess relative risk per unit dose. Two reviewers (MPL, KA) independently coded the information from the final 93 papers into a Microsoft Excel spreadsheet, and various semi-automated procedures were used to prepare the analytical database from these data; the coding was also checked by a third author (NH). Further details, including search criteria and dosimetry ascertainment, are given in supplement S1. Information coded included whether the exposure was chronic (ie, low dose rate exposure (<5 mGy/h)),<sup>28</sup> or acute (ie, at dose rates above this level).

### Classification of outcome measures

We used four major subtypes of cardiovascular disease determined a priori, which are more or less as used in an older meta-analysis,<sup>21</sup> namely:

- Ischaemic heart disease (ICD-9 410-414, ICD-10 I20-I25);
- Heart disease apart from ischaemic heart disease (ICD-9 390-398, 402, 404, 415-429, ICD-10 I11, I13, I26-I52);

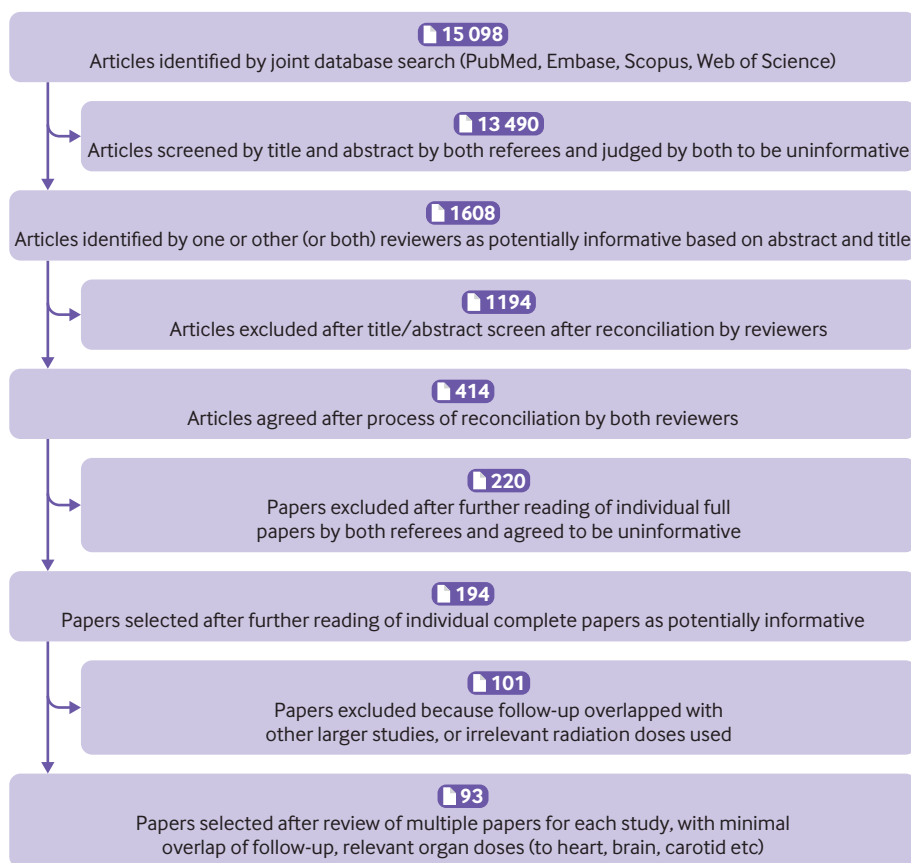


Fig 1 | PRISMA flow diagram showing exclusions made to derive the final set of studies used

- Cerebrovascular disease (ICD-9 430-438, ICD-10 I60-I69); and
- All other types of cardiovascular disease (ICD-9 399-401, 403, 405-409, 439-459, ICD-10 I00-I10, I12, I14-I19, I53-59, I70-I99).

As well as these four outcome categories, we also sought information on cardiovascular disease overall. We map study endpoints to these four endpoint groups and to the endpoint for all cardiovascular disease in supplement S3 tables S3.1 and S3.2.

### Statistical methods

The basis of all estimations of radiation risk is the value of excess relative risk per unit of effective dose (excess relative risk per Sv) or absorbed dose (excess relative risk per Gy) of radiation exposure. The excess relative risk (ERR) is related to the relative risk (RR) by  $ERR=RR-1$ , so that the excess relative risk per unit dose is  $ERR/D=(RR-1)/D$  where D represents dose. For absorbed dose, most publications use unweighted absorbed dose (Gy), but some use weighted absorbed dose (Gy) to account for the higher biological effectiveness of neutrons compared with photons (eg, in the Life Span Study<sup>20</sup>). The basis for this use in most studies is fitting of a model in which the disease or death rate (cases or deaths per year) in the group with age *a*, sex *s*, organ/tissue absorbed dose D (in Gy) is given by:  $\lambda(a,s) \times [1 + \alpha D]$  for some function  $\lambda(a,s)$

representing the disease/death rate without radiation exposure. The parameter  $\alpha$  is the excess relative risk per Gy. We collected additional information on maximum radiation dose, maximum radiation dose rate, age at exposure (a grouped variable), mortality versus incidence for each study endpoint within each study. All survivors of Hodgkin's lymphoma were deemed young adults for the purposes of analysis by age at exposure. All survivors of other cancers, except when these were treated in childhood, were assumed to be treated in adulthood.

An aggregate estimate of excess relative risk per Gy is computed across subsets of these studies by use of random effects models, using standard statistical methods (ie, meta-regression). In certain fits, adjustment was made for specific factors (ie, dose rate, level of maximum dose, mortality v morbidity, age group, disease endpoint), but in other fits, the main effect of the specific factor was assessed. Random effects models are fitted by restricted maximum likelihood because of the theoretically superior performance, in particular the absence of bias in the estimates of variance.<sup>29 30</sup> Ordinary maximum likelihood fits were also used because these facilitate comparison of nested models; in particular, to test against improvement over the null (ie, lack of homogeneity of risk), where homogeneity of risk across categories is the assumed null hypothesis. Random effects models

**Table 1 | Studies considered in systematic review and meta-analysis**

Endpoint	No. of endpoints within studies	No. of studies
<b>All studies</b>		
Ischaemic heart disease	38	38
Other heart	27	16
Cerebrovascular disease	32	31
Other cardiovascular disease	12	10
All cardiovascular disease*	105	86
All studies/endpoints†	157	93
<b>Studies with mean bias score <math>\geq 4</math></b>		
Ischaemic heart disease	17	17
Other heart	13	7
Cerebrovascular disease	14	14
Other cardiovascular disease	8	7
All cardiovascular disease*	48	34
All studies/endpoints†	75	34
<b>Studies with mean quality score <math>\geq 4</math></b>		
Ischaemic heart disease	4	4
Other heart	1	1
Cerebrovascular disease	4	4
Other cardiovascular disease	4	4
All cardiovascular disease*	10	9
All studies/endpoints†	15	9

\*Considering maximal set of all non-overlapping endpoints within a study.

†Considering maximal set of all non-overlapping endpoints within a study, as well as all non-overlapping endpoints within each of the four specific cardiovascular disease subtypes (ischaemic heart disease, other heart, cerebrovascular disease, other cardiovascular disease).

were also fitted by use of the one step approximation of DerSimonian and Laird.<sup>31</sup> Residual heterogeneity was assessed using Cochran's Q statistic. The  $I^2$  statistic of Higgins and Thompson<sup>32</sup> was computed to assess the contribution of heterogeneity to the aggregate data. These results are expressed as a percentage, so that a value near 0% implies little estimated interstudy heterogeneity relative to the intrastudy variance, and values near 100% that the interstudy heterogeneity dominates the intrastudy variance.<sup>32</sup> Confidence intervals on the  $I^2$  statistic were derived by use of the method of Knapp and Hartung.<sup>33</sup>

To assess selection or publication bias, funnel plots were used. Funnel plots are scatterplots of the central estimates (here, of excess relative risk) against estimates of standard error, and as discussed by Egger and colleagues,<sup>34 35</sup> are useful qualitative means of detecting various types of selection bias, in particular publication bias. If the funnel plot has the form of an inverted symmetrical funnel, then selection bias is considered to be unlikely.<sup>34 35</sup> More formal tests of selection or publication bias were also conducted using the test statistic suggested by Egger and colleagues.<sup>34</sup> We also used the trim-and-fill method of Duval and Tweedie<sup>36</sup> to assess the likely extent of the change in excess relative risk that can result from selection bias.

The Risk Of Bias In Non-randomised Studies-of Interventions (ROBINS-I) framework<sup>37</sup> was used to assess risk of bias associated with various characteristics of each study. A separate and objectively defined study quality score was also determined (supplement S1). Both scores were used to exclude lower quality studies.

All statistical models were fitted using the metafor package<sup>38 39</sup> in R.<sup>40</sup> Forest plots were prepared by use

of the forestplot package<sup>41</sup> in R.<sup>40</sup> Results of the meta-analysis were based on the data given in supplement S5.

### Estimates of population risks

We used pooled excess relative risk from the meta-analysis to derive population based excess absolute risk estimates according to underlying cause specific mortality rates for each population. Specifically, we used estimates for England and Wales for 2003<sup>42</sup> and 2021,<sup>43</sup> 2021 for Japan,<sup>44</sup> 2017 for France,<sup>45</sup> 2020 for Germany,<sup>46</sup> 2020 for USA,<sup>47</sup> and 2005-09 for Canada.<sup>48</sup> We assumed a five year minimum latency period, after which the excess relative risk was assumed to apply for the remainder of life. For all of these countries listed, we estimated the risk of exposure induced death per Sv, by applying methods previously used to derive comparable estimates for radiation induced cancer.<sup>49</sup>

### Patient and public involvement statement

Two cancer survivors, Josh Mailman and Jacob Adams, both of whom received radiotherapy and diagnostic radiation doses, were consulted about the implications of the findings for diagnostic and therapeutic doses received. Plain language messages about the results will be shared with the appropriate offices at the National Cancer Institute that engage with members of the public (eg, via social media feeds) and advocacy groups. A link to the open access article will be posted on LinkedIn and ResearchGate. Some of the lead researchers (MPL, NH, LBZ) will also contact various advisory and regulatory bodies (International Commission on Radiological Protection, National Council on Radiation Protection and Measurements, United Nations Scientific Committee on the Effects of Atomic Radiation), which they are already part of. Organisations like the National Cancer Institute rely on publications like ours to inform their patient facing materials on health information websites such as cancer.gov.

### Results

MPL and NH selected 194 articles from the second stage of the systematic review by consensus. These articles were subject to a more rigorous reading, and they were removed if uninformative or overlapped too much (for more details see supplement S1) with other studies, leaving 93 articles in the final selection (fig 1).

We provide a detailed survey of the risks given in each study in supplement S3 tables S4-S6 and in supplement S4. A measure of concordance is noted between the magnitude of excess risk (excess relative risk per Gy) in many different types of study and in medically (therapeutically or diagnostically) exposed groups (supplement S3 table S3.4), or in people exposed to lower levels of radiation (supplement S3 table S3.5). A few therapeutic studies used alternative measures of dose rather than the canonical ones (supplement S3 table S3.6); these alternative dose metrics generally do not suggest very different magnitudes of risk.

**Table 2 | Meta-analysis of excess relative risk for cardiovascular diseases as a result of radiation exposure, by disease endpoint**

	No. of study endpoints	Meta excess relative risk/ Gy (95% CI)	P value for heterogeneity	P value residual heterogeneity	I <sup>2</sup> (%) (Hartung-Knapp 95% CI)
<b>Adjusting for endpoint</b>					
Ischaemic heart disease	38	0.110 (0.053 to 0.167)	0.02	<0.001	80.00 (73.04 to 93.19)
Other heart disease*	27	0.054 (-0.006 to 0.113)	—	—	—
Cerebrovascular disease	32	0.176 (0.109 to 0.244)	—	—	—
Other cardiovascular disease†	12	0.183 (0.074 to 0.292)	—	—	—
<b>Adjusting for radiation type</b>					
Low dose rate‡	55	0.223 (0.157 to 0.290)	<0.001	<0.001	76.97 (70.31 to 93.07)
Acute moderate/high dose rate§	12	0.143 (0.063 to 0.223)	—	—	—
Acute fractionated moderate/high dose rate¶	42	0.068 (0.030 to 0.106)	—	—	—
<b>Adjusting for maximum dose</b>					
Maximum dose ≤0.5 Gy	16	0.311 (0.081 to 0.540)	0.001	<0.001	70.42 (60.84 to 92.14)
Maximum dose >0.5 Gy to ≤ 1 Gy	9	0.284 (0.071 to 0.497)	—	—	—
Maximum dose >1 Gy to ≤5 Gy	23	0.159 (0.097 to 0.221)	—	—	—
Maximum dose >5 Gy	53	0.064 (0.032 to 0.096)	—	—	—
<b>Adjusting for incidence v mortality</b>					
Mortality	55	0.199 (0.134 to 0.264)	0.002	<0.001	82.17 (74.77 to 93.32)
Incidence	54	0.086 (0.047 to 0.125)	—	—	—
<b>Adjusting for age at exposure group</b>					
Childhood and in utero	18	0.083 (0.012 to 0.154)	0.23	<0.001	84.30 (77.54 to 93.97)
Young adult	4	0.059 (-0.079 to 0.197)	—	—	—
Older adult and all ages	87	0.138 (0.095 to 0.182)	—	—	—
<b>Endpoint analysis (simultaneously adjusted for low dose rate, dose ≤0.5 Gy, mortality)</b>					
Ischaemic heart disease	34	0.131 (-0.019 to 0.281)	<0.001	<0.001	55.16 (46.74 to 91.49)
Other heart disease* <sup>e</sup>	23	0.121 (-0.034 to 0.277)	—	—	—
Cerebrovascular disease	32	0.170 (0.019 to 0.321)	—	—	—
Other cardiovascular disease†	12	0.203 (0.081 to 0.325)	—	—	—

All four main endpoints are analysed together. Values used in the analysis are from supplement S3 tables S3.4-S3.5. Random effects models are fitted via restricted maximum likelihood. CI=confidence interval; acute=all people exposed at moderate or high dose rate (discussed further in the methods); young adult=all survivors of Hodgkin lymphoma (discussed further in the methods).

\*Heart disease other than ischaemic heart disease.

†Cardiovascular disease other than heart disease and cerebrovascular disease.

‡Maximum dose rate <5 mGy/h.

§Maximum dose rate ≥5 mGy/h given in single acute dose.

¶Maximum dose rate ≥5 mGy/h given in multiple fractions.

### Risk modifying factors

In the Life Span Study, excess relative risk associated with radiation for all cardiovascular disease decreases with increasing age at exposure<sup>21</sup> and borderline significant trends decrease with attained age.<sup>20 21</sup> However, for heart disease in this cohort, risk did not substantially vary by sex or time since exposure;<sup>20 21</sup> no modifying effects were reported of sex, age at exposure, or attained age.<sup>50</sup> Trends of increasing risk with increasing time since exposure were observed for some non-ischaemic heart disease endpoints in the Life Span Study<sup>50</sup> and for heart disease in UK nuclear workers<sup>51</sup> (although not for cerebrovascular disease in this group<sup>52</sup>), and for all cardiovascular disease in the International Agency for Research on Cancer 15-country study.<sup>53</sup> However, decreasing trends (for all cardiovascular disease) were documented in people given diagnostic x ray exposures as part of the treatment for tuberculosis<sup>54</sup> and in a cohort of people treated for peptic ulcers.<sup>13</sup> Patients with peptic ulcers also had slightly (albeit not significantly) lower excess relative risk for women than for men for all cardiovascular disease, ischaemic heart disease, and cardiovascular disease excluding ischaemic heart disease and cerebrovascular disease. However, excess relative risk for women was higher (but not significantly so) than that for men for cerebrovascular disease in this cohort.<sup>13</sup> In the INWORKS study, women had

significantly higher excess relative risk than did men.<sup>55</sup> Age at treatment in the Nordic study<sup>56</sup> had no effect but radiation risk with increasing age at treatment had a borderline significant reduction in a Dutch study of breast cancer survivors.<sup>57</sup>

Few studies assessed the possible modifying effect of lifestyle, medical, and environmental variables, particularly the major risk factors for cardiovascular disease (eg, smoking, obesity, diabetes, hypertension, and hypercholesterolaemia). The Nordic and Dutch studies on breast cancer,<sup>56-58</sup> and the three Dutch studies on Hodgkin lymphoma,<sup>59-61</sup> all of which have particularly rich data of this sort (supplement S3 table S3.4), found little evidence of modification of radiation dose response for various heart related endpoints for any of these variables. No modifying effects were reported of smoking or alcohol consumption in a cohort with peptic ulcers.<sup>13</sup> Recorded data for lifestyle is extensive in the Life Span Study study<sup>20</sup> (supplement S3 table S3.5) but no assessment was made of possible modifications in radiation dose-response associated with any of these risk factors. A case-control study of industrial workers at two UK nuclear plants (Sellafield and Springfield) collected information on numerous lifestyle and environmental risk factors (ie, body mass index, smoking status, diastolic and systolic blood pressure, shift work),<sup>62</sup> as did a case-control study of French nuclear fuel cycle workers (supplement

**Table 3 | Meta regression analyses of excess relative risk for each major cardiovascular disease endpoint group, with restriction by dose and dose rate**

Cardiovascular disease endpoint	No. of study endpoints	Meta excess relative risk/ Gy (95% CI)	Residual heterogeneity P value	I <sup>2</sup> (%) (Knapp-Hartung 95% CI)	Egger selection test P value	Duval-Tweedie trim-fill selection bias corrected meta excess relative risk/ Gy (95% CI)
<b>Full data</b>						
Ischaemic heart disease	38	0.073 (0.047 to 0.099)	0.01	17.80 (3.52 to 95.64)	0.002	0.073 (0.052 to 0.095)
Other heart disease*	27	0.034 (0.020 to 0.049)	0.49	0.00 (0.00 to 84.81)	0.13	0.034 (0.021 to 0.048)
Cerebrovascular disease	32	0.188 (0.093 to 0.283)	<0.001	83.56 (68.52 to 99.22)	0.05	0.184 (0.099 to 0.269)
Other cardiovascular disease†	12	0.172 (-0.029 to 0.373)	<0.001	90.27 (68.02 to 98.43)	0.49	0.172 (-0.029 to 0.373)
All cardiovascular disease (using maximal cardiovascular disease data per study)	105	0.106 (0.076 to 0.135)	<0.001	88.61 (85.58 to 95.99)	<0.001	0.102 (0.075 to 0.129)
<b>Maximum dose ≤0.5 Gy only</b>						
Ischaemic heart disease	6	0.438 (-0.131 to 1.007)	0.17	21.89 (0.00 to 99.74)	0.68	0.438 (-0.131 to 1.007)
Other heart disease*	2	-0.108 (-0.528 to 0.313)	0.76	0.00 (0.00 to 68.13)	0.39	-0.188 (-0.617 to 0.242)
Cerebrovascular disease	7	0.542 (-0.281 to 1.366)	0.24	0.00 (0.00 to >97.54)	0.79	0.542 (-0.281 to 1.366)
Other cardiovascular disease†	1	-1.80 (-11.95 to 8.35)	NA	NA	NA	NA
All cardiovascular disease (using maximal cardiovascular disease data per study)	9	0.452 (0.064 to 0.840)	0.36	17.88 (0.00 to 98.38)	0.26	0.452 (0.092 to 0.811)
<b>Low dose rate data only</b>						
Ischaemic heart disease	22	0.202 (0.085 to 0.319)	0.13	45.34 (0.00 to 86.89)	0.86	0.202 (0.085 to 0.319)
Other heart disease*	6	-0.207 (-0.456 to 0.042)	0.98	0.00 (0 to >0)	0.55	-0.241 (-0.727 to 0.246)
Cerebrovascular disease	21	0.298 (0.101 to 0.495)	<0.001	61.07 (29.38 to 99.94)	0.18	0.294 (0.130 to 0.458)
Other cardiovascular disease†	6	0.166 (-0.069 to 0.401)	0.17	39.95 (0.00 to 99.58)	0.79	0.166 (-0.069 to 0.401)
All cardiovascular disease (using maximal cardiovascular disease data per study)	41	0.229 (0.136 to 0.322)	<0.001	68.23 (31.40 to 92.81)	0.09	0.224 (0.134 to 0.315)
<b>Maximum dose ≤0.5 Gy or low dose rate only</b>						
Ischaemic heart disease	24	0.205 (0.092 to 0.318)	0.07	39.39 (0.00 to 93.39)	0.31	0.205 (0.092 to 0.318)
Other heart disease*	8	-0.168 (-0.429 to 0.094)	0.90	0.00 (0.00 to 55.12)	0.29	-0.233 (-0.567 to 0.101)
Cerebrovascular disease	23	0.306 (0.127 to 0.485)	<0.001	57.05 (22.37 to 99.87)	0.17	0.303 (0.148 to 0.458)
Other cardiovascular disease†	6	0.166 (-0.069 to 0.401)	0.17	39.95 (0.00 to 99.58)	0.79	0.166 (-0.069 to 0.401)
All cardiovascular disease (using maximal cardiovascular disease data per study)	44	0.231 (0.141 to 0.320)	<0.001	65.03 (31.27 to 94.25)	0.04	0.226 (0.141 to 0.311)

All four main endpoints and all cardiovascular disease are analysed separately. Values used in the analysis are from supplement S3 tables S3.4-S3.5. Random effects models are fitted via restricted maximum likelihood. CI=confidence interval; NA=not available.

\*Heart disease other than ischaemic heart disease.

†Cardiovascular disease other than heart disease and cerebrovascular disease.

S3 table S3.5).<sup>63</sup> Neither study assessed modifying effects on the radiation dose response, possibly because adjustments to the background risk did not change risk estimates. Information about lifestyle and environmental risk factors is available from the workers from the Mayak nuclear facility in Russia (supplement S3 table S3.5); however, analyses did not report modifying effects of these variables on the dose-response, again, possibly because adjustments to the baseline risk had little effect.<sup>64-68</sup>

Only modest information is available of the modifying effects of cardiotoxic treatment on radiation response. No modifying effects were reported in the Nordic and Dutch breast cancer case-control studies,<sup>56-58</sup> nor were such modifications indicated in Dutch Hodgkin lymphoma studies.<sup>59-61</sup>

### Results of meta-analysis

Table 1, table 2, table 3, table 4, table 5, and figure 2 report the results of the meta-analysis. We found (table 3) that radiation exposure was associated with a generally significant meta-excess relative risk per Gy for all cardiovascular disease (0.11 (95% confidence interval 0.08 to 0.14)), ischaemic heart

disease (0.07 (0.05 to 0.10)), other heart disease (0.03 (0.02 to 0.05)), cerebrovascular disease (0.19 (0.09 to 0.28)), and other cardiovascular disease (0.17 (-0.03 to 0.37)). Meta excess relative risk per Gy varied significantly between subtypes of disease (table 2). For all cardiovascular disease, a significant meta excess relative risk per Gy was also noted for all levels of maximum radiation dose, even for lower-dose exposures with maximum exposure of 0.5 Gy or less (0.45 per Gy (95% confidence interval 0.06 to 0.84)), and if combined with low dose rate studies, this was also the case for ischaemic heart disease (0.20 (95% confidence interval 0.09 to 0.32); table 3). Significant heterogeneity was noted in relation to maximum radiation dose (P=0.001), with cardiovascular disease risk higher when maximum radiation dose was at 0.5 Gy or less, 0.5-1 Gy, and 1-5 Gy than at more than 5 Gy (table 2). We also observed significant difference in meta excess relative risk per Gy by radiation dose rate (P<0.001), with risk appreciably higher for low dose rate exposure. Table 2 also suggests that risks are significantly higher (P=0.002) for mortality endpoints compared with those of incidence.

**Table 4 | Meta-analysis of higher quality estimates of excess relative risk for cardiovascular diseases as a result of radiation exposure, by disease endpoint. All four main endpoints are analysed together**

	No of study endpoints	Excess relative risk (meta excess relative risk/Gy) (95% CI)	P value for heterogeneity	P value residual heterogeneity	I <sup>2</sup> (%) (Hartung-Knapp 95% CI)
<b>Analysis using mean bias score ≥4</b>					
Adjusting for endpoint:					
Ischaemic heart disease	17	0.132 (0.027 to 0.237)	0.50	<0.001	72.16 (40.34 to 85.62)
Other heart disease*	13	0.117 (-0.032 to 0.266)	—	—	—
Cerebrovascular disease	14	0.236 (0.102 to 0.370)	—	—	—
Other cardiovascular disease†	8	0.208 (0.057 to 0.359)	—	—	—
Adjusting for radiation type					
Low dose rate‡	26	0.250 (0.132 to 0.369)	0.25	<0.001	72.92 (41.84 to 86.38)
Acute moderate/high dose rate§	12	0.148 (0.048 to 0.248)	—	—	—
Acute fractionated moderate/high dose rate¶	14	0.117 (0.005 to 0.229)	—	—	—
Adjusting for maximum dose:					
Maximum dose ≤0.5 Gy	13	0.303 (0.067 to 0.538)	0.55	<0.001	72.97 (36.47 to 87.89)
Maximum dose >0.5 Gy, ≤1 Gy	4	0.158 (-0.163 to 0.479)	—	—	—
Maximum dose >1 Gy, ≤5 Gy	18	0.149 (0.059 to 0.239)	—	—	—
Maximum dose >5 Gy	11	0.106 (-0.026 to 0.238)	—	—	—
Adjusting for incidence v mortality:					
Mortality	34	0.214 (0.120 to 0.308)	0.20	<0.001	76.10 (47.80 to 86.95)
Incidence	18	0.129 (0.042 to 0.215)	—	—	—
<b>Adjusting for age at exposure group</b>					
Childhood and in utero	5	0.140 (-0.064 to 0.344)	0.94	<0.001	77.57 (50.29 to 88.28)
Young adult	1	0.141 (-0.417 to 0.699)	—	—	—
Older adult and all age	46	0.172 (0.102 to 0.243)	—	—	—
<b>Endpoint analysis (simultaneously adjusted for low dose rate, dose ≤0.5 Gy, mortality)</b>					
Ischaemic heart disease	15	0.397 (0.047 to 0.747)	0.16	<0.001	56.57 (12.64 to 81.54)
Other heart disease*	9	0.474 (0.057 to 0.891)	—	—	—
Cerebrovascular disease	14	0.539 (0.157 to 0.921)	—	—	—
Other cardiovascular disease†	8	0.595 (0.175 to 1.015)	—	—	—
<b>Analysis using mean quality score ≥4</b>					
Adjusting for endpoint:					
Ischaemic heart disease	4	0.109 (-0.236 to 0.453)	0.47	<0.001	92.72 (63.27 to 98.79)
Other heart disease*	1	0.038 (-0.451 to 0.526)	—	—	—
Cerebrovascular disease	4	0.290 (-0.057 to 0.636)	—	—	—
Other cardiovascular disease†	4	0.270 (-0.036 to 0.575)	—	—	—
Adjusting for radiation type					
Low dose rate‡	8	0.299 (0.039 to 0.558)	0.36	<0.001	91.15 (65.49 to 97.93)
Acute high dose rate§	3	0.211 (-0.060 to 0.482)	—	—	—
Acute fractionated high dose rate¶	2	0.056 (-0.253 to 0.365)	—	—	—
Adjusting for maximum dose					
Maximum dose ≤0.5 Gy	5	0.364 (-8.955 to 9.684)	0.72	<0.001	93.23 (63.08 to 99.98)
Maximum dose >0.5 Gy, ≤1 Gy	0	NA	—	—	—
Maximum dose >1 Gy, ≤5 Gy	3	0.203 (-0.138 to 0.544)	—	—	—
Maximum dose >5 Gy	4	0.056 (-0.340 to 0.452)	—	—	—
Adjusting for incidence v mortality					
Mortality	7	0.342 (0.043 to 0.641)	0.26	<0.001	92.34 (73.54 to 97.75)
Incidence	6	0.154 (-0.023 to 0.332)	—	—	—
<b>Adjusting for age at exposure group</b>					
Childhood and in utero	0	NA	0.35	<0.001	92.50 (73.22 to 97.93)
Young adult	1	0.038 (-0.390 to 0.466)	—	—	—
Older adult and all age	12	0.229 (0.060 to 0.397)	—	—	—
<b>Endpoint analysis (simultaneously adjusted for low dose rate, dose ≤0.5 Gy, mortality)</b>					
Ischaemic heart disease	3	6.84 (-19.03 to 32.72)	0.003	0.67	0.00 (0.00 to >27.08)
Other heart disease*	1	6.81 (-19.07 to 32.68)	—	—	—
Cerebrovascular disease	4	-1.36 (-12.85 to 10.14)	—	—	—
Other cardiovascular disease†	4	-0.90 (-12.39 to 10.60)	—	—	—

Values for the analysis are from supplement S3 tables S3.4-S3.5. Analysis is restricted to studies with mean bias score ≥4 or mean quality score ≥4 (as given in supplement S3 tables S3.1, S3.2). Random effects models are fitted via restricted maximum likelihood. CI=confidence interval; NA=not available.

\*Heart disease other than ischaemic heart disease.

†Cardiovascular disease other than heart disease and cerebrovascular disease.

‡Maximum dose rate <5 mGy/h.

§Maximum dose rate ≥5 mGy/h given in single acute dose.

¶Maximum dose rate ≥5 mGy/h given in multiple fractions.

Table 3 suggests that for ischaemic heart disease, cerebrovascular disease, and all cardiovascular disease, meta excess relative risk per Gy was

significantly elevated for low dose rate data or low dose rate combined with studies in which maximum dose was 0.5 Gy or less. For these three endpoints, the meta

**Table 5 | Meta regression analyses of higher quality estimates of excess relative risk for each major cardiovascular disease endpoint group, with restriction by dose and dose rate**

	No. of study endpoints	Excess relative risk (meta excess relative risk/Gy) (95% CI)	Residual heterogeneity P value	I <sup>2</sup> (%) (Knapp-Hartung 95% CI)	Egger selection test P value	Duval-Tweedie trim-fill selection bias corrected meta excess relative risk/Gy (95% CI)
<b>Analysis using mean bias score ≥4</b>						
Full data:						
Ischaemic heart disease	17	0.099 (0.059 to 0.139)	0.17	0.37 (0.00 to 98.72)	0.13	0.099 (0.066 to 0.132)
Other heart disease*	13	0.108 (-0.007 to 0.224)	0.25	36.00 (0.00 to 74.22)	0.56	0.108 (-0.007 to 0.224)
Cerebrovascular disease	14	0.237 (0.096 to 0.378)	<0.001	72.86 (21.64 to 94.48)	0.15	0.214 (0.074 to 0.353)
Other cardiovascular disease†	8	0.203 (-0.067 to 0.474)	<0.001	86.31 (56.18 to 98.15)	0.82	0.203 (-0.067 to 0.474)
All cardiovascular disease (using maximal cardiovascular disease data per study)	48	0.148 (0.087 to 0.209)	<0.001	85.19 (72.47 to 94.61)	0.007	0.136 (0.077 to 0.195)
Maximum dose ≤0.5 Gy or low dose rate:						
Ischaemic heart disease	12	0.153 (0.080 to 0.227)	0.23	0.00 (0.00 to 99.53)	0.28	0.153 (0.095 to 0.211)
Other heart disease*	7	-0.107 (-0.471 to 0.257)	0.85	0.00 (0.00 to 45.34)	0.38	-0.214 (-0.637 to 0.209)
Cerebrovascular disease	11	0.389 (0.191 to 0.588)	0.52	14.48 (0.00 to 81.22)	0.29	0.377 (0.171 to 0.584)
Other cardiovascular disease†	2	0.299 (-0.226 to 0.825)	0.69	0.00 (0.00 to >88.17)	NA	NA
All cardiovascular disease (using maximal cardiovascular disease data per study)	22	0.214 (0.090 to 0.339)	<0.001	72.09 (34.23 to 96.55)	0.09	0.211 (0.095 to 0.327)
<b>Analysis using mean quality score ≥4</b>						
Full data:						
Ischaemic heart disease	4	0.105 (0.025 to 0.186)	0.39	27.16 (0.00 to >99.99)	0.63	0.105 (0.041 to 0.170)
Other heart disease*	1	0.038 (-0.035 to 0.111)	NA	NA	NA	0.108 (-0.007 to 0.224)
Cerebrovascular disease	4	0.289 (-0.147 to 0.726)	<0.001	91.14 (48.40 to >99.99)	0.50	0.288 (-0.045 to 0.622)
Other cardiovascular disease†	4	0.254 (-0.326 to 0.835)	<0.001	88.16 (47.23 to 99.52)	0.56	0.258 (-0.157 to 0.673)
All cardiovascular disease (using maximal cardiovascular disease data per study)	10	0.204 (0.039 to 0.369)	<0.001	94.71 (84.55 to 98.28)	0.97	0.204 (0.039 to 0.369)
Maximum dose ≤0.5 Gy or low dose rate:						
Ischaemic heart disease	3	0.140 (0.050 to 0.230)	0.67	0.00 (0.00 to >43.68)	0.65	0.140 (0.050 to 0.230)
Other heart disease*	0	NA	NA	NA	NA	NA
Cerebrovascular disease	3	0.460 (0.317 to 0.603)	0.62	0.00 (0.00 to >16.79)	0.52	0.460 (0.317 to 0.603)
Other cardiovascular disease†	2	0.299 (-0.226 to 0.825)	0.69	0.00 (0.00 to >88.17)	NA	NA
All cardiovascular disease (using maximal cardiovascular disease data per study)	5	0.298 (0.105 to 0.491)	<0.001	82.50 (34.55 to 95.70)	0.78	0.297 (0.101 to 0.494)

All four main endpoints and all cardiovascular disease are analysed separately. Analysis is restricted to studies with mean bias score ≥4 or mean quality score ≥4 (as given in Supplement S3 tables S3.1, S3.2). Random effects models are fitted via restricted maximum likelihood. CI=confidence interval; NA=not available.  
 \*Heart disease other than ischaemic heart disease.  
 †Cardiovascular disease other than heart disease and cerebrovascular disease.

excess relative risk per Gy are higher for maximum dose under 0.5 Gy or for low dose rate (separately or together), and the residual heterogeneity (as measured by the I<sup>2</sup> statistic) tended also to be lower.

The funnel plots given in figure 2 do not suggest any material selection or publication bias, and more formal tests of selection suggest presence of such bias for only a few endpoints (eg, ischaemic heart disease using all data (table 3)). The Duval-Tweedie trim-and-fill bias corrected meta excess relative risk per Gy and confidence interval generally differ little from the uncorrected estimates (table 3). Interstudy heterogeneity is substantial, particularly for cerebrovascular disease and all cardiovascular disease, with values of the I<sup>2</sup> statistic generally above 50%.

When analysis was restricted to higher quality studies, with a mean bias score of at least 4 or a mean study quality score of at least 4, results were similar (table 4, table 5). A notable feature of the higher quality studies is the weakening evidence of interstudy heterogeneity, which remains significant only (in some cases) for cerebrovascular disease, all other cardiovascular disease, and all cardiovascular disease (table 5). However, with higher quality studies, only the meta excess relative risk per Gy for all cardiovascular disease and the subtypes ischaemic heart disease and cerebrovascular disease generally remained statistically significant; and heterogeneity by maximum dose, dose rate, and mortality versus incidence are no longer significant (P>0.2) (table 4, table 5). Nevertheless, indications suggest that risk is



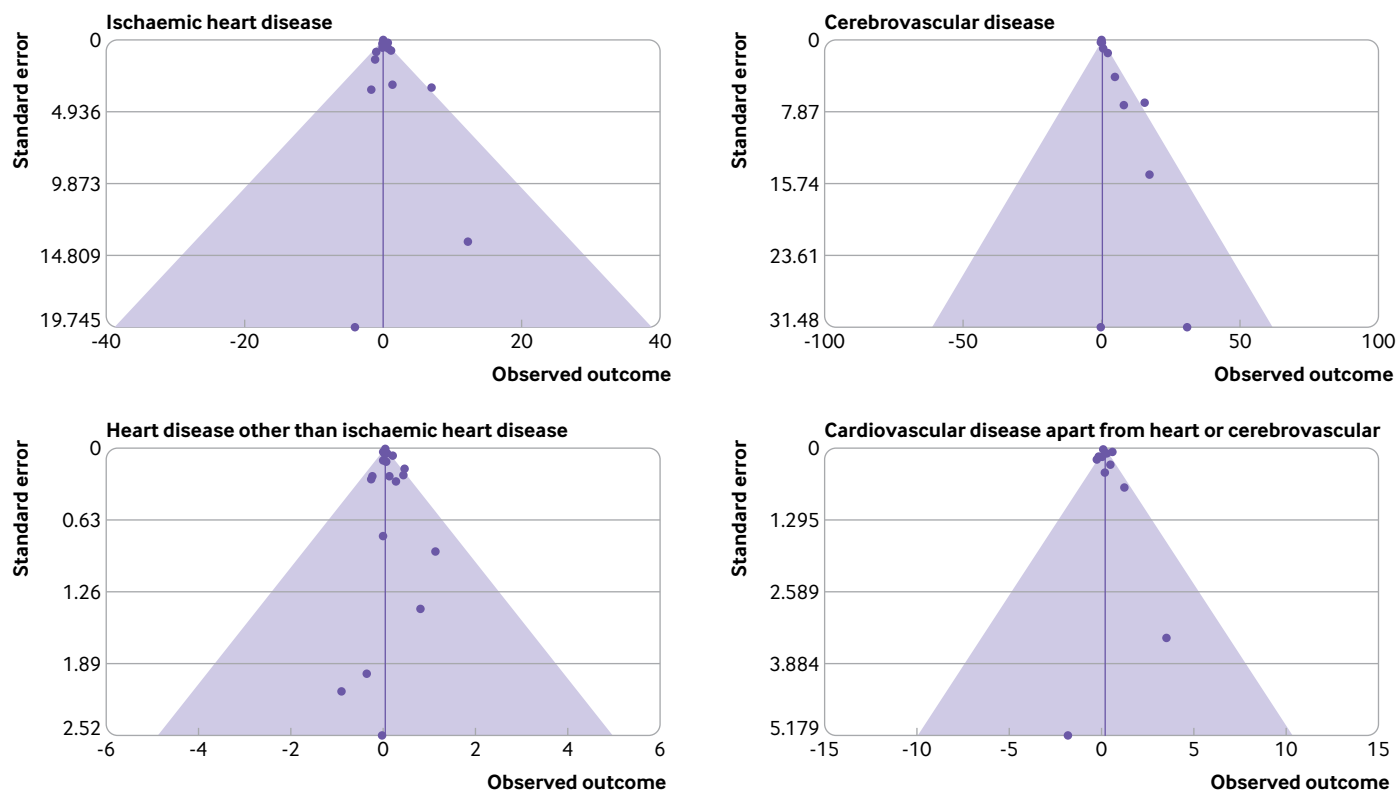


Fig 2 | Funnel plot of risks by four major cardiovascular disease endpoints. A study without appreciable selection bias should have a more or less balanced funnel plot, with points spread more or less equally to left and right of the vertical axis of the funnel

higher at lower doses, and at a lower dose rate (table 4, table 5). We saw a slight tendency for meta excess relative risk per Gy to be higher (although also more uncertain) among higher quality studies, for all four main endpoints (supplement S3 figures S3.1, S3.2).

The alternative fitting methods (maximum likelihood, DerSimonian-Laird one step) yielded very similar estimates and confidence intervals as those of restricted maximum likelihood, the default method used (supplement S3 table S3.3).

Sensitivity analyses in which we excluded Mayak morbidity data, Mayak mortality data or the Canadian data of Zielinski and colleagues<sup>69</sup>, or in which we added the Los Alamos<sup>70</sup> and Rochester thymus<sup>71</sup> data to the analysis did not suggest large changes. The most substantial change resulted from removal of the Mayak incidence data, when the meta excess relative risk per Gy for all cardiovascular disease reduced from 0.11 (95% confidence interval 0.08 to 0.14) to 0.09 (95% confidence interval 0.06 to 0.11), but hardly changed when the Mayak mortality data or the Canadian data were removed, or the Los Alamos and Rochester cohorts were added (supplement S3 table S3.7).

#### Population risks

Population based excess absolute risk estimates for radiation exposure induced death for all cardiovascular disease range from 2.33% per Gy (95% confidence interval 1.69% to 2.38%) for England and Wales (using 2021 rates) to 3.66% per Gy (2.65% to 4.68%)

for Germany, largely reflecting the underlying risk of cardiovascular disease mortality (table 6). Estimated mortality risks of cardiovascular disease are generally dominated by cerebrovascular disease, with the next largest contribution from ischaemic heart disease (table 6). If the 2003 England and Wales rates are used, the excess absolute risk is appreciably higher, 3.94% per Gy (2.85% to 5.03%), reflecting the much higher proportion of deaths due to cardiovascular disease in this earlier population (39.88% v 23.57%). Years of life lost per Gy range from 0.190 (0.137 to 0.243) for Japan to 0.373 (0.270 to 0.477) for USA. The markedly higher figures for years of life lost per death induced by radiation for USA compared with other populations are largely artefactual because US mortality rates are not published for any age group older than 80-84 years. Therefore, mortality rates at older ages, which very steeply increase for all causes as well as cardiovascular disease in all other populations, must be assumed to be those given by this age group. If, for example, the data for the 85-89 and 90 and older year groups for the 2021 population in England and Wales were removed, so that effectively rates at ages older than 80 years simply reflected those for ages 80-84, then all circulatory risk of exposure induced death would decrease to 2.08% per Gy (1.51% to 2.66%). Years of life lost per Gy would markedly increase to 0.311 per Gy (0.225 to 0.398), so that years life lost per radiation induced death, which is simply the quotient of these (years of life lost/risk of exposure induced death),

**Table 6 | Estimated excess absolute risk of radiation exposure induced death (% per Gy), years life lost per Gy, and years life lost per radiation induced death, and 95% confidence intervals for various subtypes of cardiovascular disease by country, using subtype specific restricted maximum likelihood estimates for all data from table 3**

Cardiovascular disease endpoint	Canada (2005-09 rates)	England and Wales (2021 rates)	England and Wales (2003 rates)	France (2017 rates)	Germany (2020 rates)	Japan (2021 rates)	USA (2020 rates)
Baseline proportion of deaths due to cardiovascular disease (%)	33.97	23.57	39.88	26.45	37.00	25.35	25.57
Risk of exposure induced deaths $\times 10^{-2}$ :							
Ischaemic heart disease	1.20 (0.77 to 1.62)	0.64 (0.41 to 0.87)	1.30 (0.84 to 1.75)	0.39 (0.25 to 0.53)	0.89 (0.58 to 1.21)	0.30 (0.19 to 0.40)	0.73 (0.48 to 0.99)
Other heart disease	0.21 (0.13 to 0.30)	0.21 (0.12 to 0.29)	0.18 (0.11 to 0.25)	0.37 (0.22 to 0.52)	0.27 (0.16 to 0.38)	0.37 (0.22 to 0.52)	0.26 (0.15 to 0.37)
Cerebrovascular disease	1.25 (0.62 to 1.88)	0.94 (0.46 to 1.41)	2.04 (1.01 to 3.07)	0.99 (0.49 to 1.50)	1.00 (0.49 to 1.51)	1.26 (0.62 to 1.90)	0.83 (0.41 to 1.25)
Other cardiovascular disease	0.47 (-0.08 to 1.02)	0.41 (-0.07 to 0.88)	0.61 (-0.10 to 1.31)	0.61 (-0.10 to 1.32)	1.61 (-0.27 to 3.49)	0.42 (-0.07 to 0.90)	0.33 (-0.06 to 0.73)
All cardiovascular disease	3.37 (2.43 to 4.30)	2.33 (1.69 to 2.98)	3.94 (2.85 to 5.03)	2.63 (1.90 to 3.35)	3.66 (2.65 to 4.68)	2.52 (1.82 to 3.22)	2.53 (1.83 to 3.23)
Years of life lost per Gy:							
Ischaemic heart disease	0.101 (0.065 to 0.136)	0.061 (0.040 to 0.083)	0.112 (0.072 to 0.151)	0.032 (0.021 to 0.044)	0.077 (0.050 to 0.105)	0.027 (0.018 to 0.037)	0.108 (0.070 to 0.146)
Other heart disease	0.016 (0.010 to 0.023)	0.016 (0.009 to 0.023)	0.013 (0.008 to 0.019)	0.024 (0.014 to 0.034)	0.021 (0.012 to 0.030)	0.024 (0.014 to 0.033)	0.039 (0.023 to 0.055)
Cerebrovascular disease	0.097 (0.048 to 0.145)	0.075 (0.037 to 0.112)	0.148 (0.073 to 0.223)	0.075 (0.037 to 0.113)	0.086 (0.043 to 0.130)	0.101 (0.050 to 0.152)	0.118 (0.058 to 0.178)
Other cardiovascular disease	0.038 (-0.006 to 0.082)	0.035 (-0.006 to 0.075)	0.051 (-0.009 to 0.111)	0.048 (-0.008 to 0.105)	0.131 (-0.022 to 0.283)	0.035 (-0.006 to 0.075)	0.051 (-0.008 to 0.110)
All cardiovascular disease	0.273 (0.197 to 0.349)	0.201 (0.145 to 0.256)	0.317 (0.229 to 0.406)	0.191 (0.138 to 0.245)	0.305 (0.221 to 0.390)	0.190 (0.137 to 0.243)	0.373 (0.270 to 0.477)
Years of life lost per radiation induced death:							
Ischaemic heart disease	8.419 (8.418 to 8.421)	9.557 (9.556 to 9.558)	8.613 (8.612 to 8.615)	8.220 (8.220 to 8.221)	8.668 (8.667 to 8.669)	9.238 (9.238 to 9.238)	14.719 (14.718 to 14.720)
Other heart disease	7.697 (7.697 to 7.698)	7.719 (7.718 to 7.719)	7.497 (7.496 to 7.497)	6.476 (6.476 to 6.477)	7.860 (7.860 to 7.860)	6.448 (6.448 to 6.448)	14.931 (14.931 to 14.932)
Cerebrovascular disease	7.744 (7.742 to 7.746)	7.950 (7.949 to 7.952)	7.261 (7.258 to 7.264)	7.544 (7.542 to 7.546)	8.598 (8.597 to 8.600)	7.986 (7.984 to 7.988)	14.295 (14.294 to 14.297)
Other cardiovascular disease	8.114 (8.112 to 8.116)	8.495 (8.493 to 8.496)	8.438 (8.436 to 8.441)	7.912 (7.910 to 7.915)	8.121 (8.116 to 8.125)	8.312 (8.311 to 8.314)	15.098 (15.097 to 15.100)
All cardiovascular disease	8.121 (8.118 to 8.124)	8.588 (8.586 to 8.591)	8.056 (8.053 to 8.060)	7.292 (7.290 to 7.295)	8.334 (8.331 to 8.336)	7.551 (7.549 to 7.554)	14.746 (14.743 to 14.748)

Calculations use a test dose of 0.1 Gy and are given to populations assumed to be in equilibrium having the overall mortality and cardiovascular disease mortality rates of the given population. The population is assumed to be followed up to age 120 years.

would nearly double, to 14.927 (14.925 to 14.929). This value is quite close to the USA figure of 14.746 (14.743 to 14.748). This highlights the effect of age groups older than 80 years on population risks.

**Discussion**

**Principal findings**

Our comprehensive meta-analysis, covering a range of individuals who had been exposed to radiation medically (therapeutically or diagnostically), occupationally, or environmentally, shows a significantly increased excess risk per unit dose for most subtypes of cardiovascular disease. We noted heterogeneity between studies, possibly resulting from variation between studies in unmeasured confounders or effect modifiers, however, this evidence of interstudy heterogeneity is markedly reduced among studies with a maximum dose of less than 0.5 Gy or low dose rate or if consideration is restricted to higher quality studies. For ischaemic heart disease, cerebrovascular disease and all cardiovascular disease, risks were larger per unit dose for lower dose (inverse dose effect) and lower dose rate and fractionated exposures (inverse dose fractionation effect). However, the evidence was weaker if attention was restricted to higher quality studies.

**Comparison with other studies**

A reduction in mean cumulative dose increased excess relative risk per unit dose, a finding that is consistent with the Life Span Study of atomic bomb survivors. In particular, analysis of Life Span Study data, considering dose error, suggests a substantial downwardly curving dose response for cardiovascular disease,<sup>72</sup> which has also been observed for ischaemic heart disease<sup>51</sup> and cerebrovascular disease<sup>52 55</sup> in workers at nuclear facilities. Our finding that increased fractionation increases cardiovascular disease risk is consistent with what was found in an analysis of the Canadian tuberculosis fluoroscopy cohort<sup>73</sup>; however, no evidence of such a fractionation effect was reported if latency was more than or less than 10 years. Additionally, no such evidence was found in the pooled analysis of the Massachusetts and Canadian data,<sup>54</sup> although the pattern of sparing and enhancing effects of dose protraction is complex depending on irradiation regimens at high dose.<sup>74</sup> Excess relative risk reduced with increasing age at exposure for stroke and all cardiovascular disease in the Life Span Study,<sup>21</sup> but not for heart disease.<sup>50</sup>

The excess relative risks that we derived (table 2, table 3) were generally consistent with those of a previous systematic review and meta-analysis, published over a decade ago, of moderate to low dose studies,<sup>21</sup> and with those of a subsequent non-systematic review.<sup>27</sup> Given the overlap in the moderate to low dose studies considered in our paper and previously, this consistency is perhaps unsurprising. However, as suggested by the results of the meta-regression analysis (table 2, table 3), even if the differences between risks at low and moderate to high

dose rate were not substantial, they were nevertheless significant.

### Population risks

Ischaemic heart disease and cerebrovascular disease are the most strongly associated with radiation, even if analysis was restricted to less than 0.5 Gy or low dose rate data (table 2). The excess absolute risk coefficients that we derived for a UK population of 2.3% (for a 2021 population) to 3.9% per Gy (for a 2003 population) (table 6) were slightly lower than, but of similar extent to, those estimated for cancer mortality by the United Nations Scientific Committee on the Effects of Atomic Radiation,<sup>49</sup> which for a 2003 UK population were in the range 4.4-5.2%. At the population level, the largest risks were for cerebrovascular disease and ischaemic heart disease (table 6). For a 2003 UK population, the years of life lost from all cardiovascular disease per Gy was about 0.3 (table 6), which is slightly lower than but similar to the estimated 0.6-0.7 years of life lost per Gy for all solid cancers, estimated by the United Nations Scientific Committee on the Effects of Atomic Radiation.<sup>49</sup> This finding has considerable implications for the system of radiological protection, assuming that the extrapolation is permissible, even, for example, over the restricted dose range 0-0.5 Gy. This added risk would nearly double the low dose detriment. Even the restricted range risks were based on people who were potentially exposed to up to 0.5 Gy of radiation, possibly augmented by the people exposed at low dose rate. This level is not what is normally thought of as low dose, which usually refers to risks at doses of less than 0.1 Gy.<sup>75</sup> The available evidence does not suggest that lifestyle, environmental, or medical risk factors appreciably modify the excess relative risk related to radiation. Nevertheless, because most major risk factors for cardiovascular disease (smoking, obesity, diabetes, hypertension, elevated cholesterol, unhealthy diet, physical inactivity, and psychosocial factors) multiply (by factors of two or more) the normal risk of cardiovascular disease,<sup>4-6</sup> <sup>76-77</sup> lifetime radiation risk might greatly increase among people with these extra risk factors. This effect should prompt extra vigilance to control modifiable cardiovascular risk factors in patients who receive substantial doses of ionising radiation as part of their medical care. The risks that we estimated for most types of radiation exposure received by the population were relatively trivial. However, this might not be the case for patients who received radiotherapy, radionuclide treatment, a fluoroscopically guided interventional procedure with high doses of radiation, or numerous medical imaging or fluoroscopically guided procedures. Patients can receive doses of 0.1 Gy to relevant organs when they receive radiotherapy for benign conditions, radionuclide treatment, multiple high dose diagnostic (computed tomography and nuclear medicine), or fluoroscopic procedures. The lifetime risks that we estimated (2.3-3.9% Gy<sup>-1</sup> (table 6)), imply that the 10-20 Gy that might be delivered to the heart with some types of treatment (supplement

S3 table S3.5) would result in lifetime cardiovascular disease risk that exceeds 50%. However, such risks are to some extent an inevitable consequence of life saving treatment, and the benefits of such therapy in general will outweigh cardiovascular disease risks. Doses from most diagnostic procedures are considerably lower, so that, for example, a typical computed tomography scan might deliver a dose of between 0.0005 and 0.015 Gy to the heart.<sup>78-79</sup> This dosage together with the risks in table 6 suggests that a group of 10 000 people in the UK each exposed to 10 procedures of this sort might expect between 0.2-13.0 excess ischaemic heart disease deaths over a lifetime.

### Mechanistic information and relevant target tissue

Various reviews suggest candidate biological mechanisms.<sup>25 80-82</sup> Inflammatory mechanisms are plausible, if not completely understood, means by which high doses of radiation could affect the cardiovascular system<sup>81</sup>; radiation effects on the immune system might also play a part.<sup>83</sup> At lower doses, much less is known. Numerous mechanisms have been proposed, for example, monocyte cell killing in the arterial intima,<sup>84</sup> likewise, radiation induced endothelial cell senescence and associated monocyte adhesion<sup>85-89</sup>; however, these mechanisms remain speculative.

Evidence from the radiotherapy cohorts suggests that radiation dose to the heart could be the most relevant for ischaemic heart disease.<sup>56</sup> Doses to the heart and thyroid (surrogate for a carotid dose) might also be relevant for cerebrovascular disease; however, doses to the brain are unlikely to be associated.<sup>13</sup> The generally uniform whole body radiation with low linear energy transfer in the lower dose cohorts is uninformative as to specific target tissues. In many occupational studies, effective dose is used, in which absorbed dose to each organ is weighted by appropriate tissue weighting factors; this contrasts with the absorbed organ dose that is used elsewhere. However, these different dose metrics would not be expected to be markedly different for the penetrating ionising radiations with low linear energy transfer considered here, so would not substantially contribute to heterogeneity in radiation risk. The consistency of risks, across a wide range of doses (supplement S3 tables S3.4, S3.5) suggests that target tissues and associated mechanisms might be the same for all levels of dose.

### Limitations

A concerning feature of our meta-analysis is that for many endpoints, and in particular for cerebrovascular disease, heterogeneity was significant ( $P < 0.001$ ), and this together with high values (generally  $> 50\%$ ) of the  $I^2$  statistic imply that a material proportion of the variance was due to interstudy heterogeneity (table 3). This issue makes interpretation of summary measures of risk problematic. However, when we restricted analysis to lower dose or dose rate studies (table 3) or when we restricted attention to the higher quality studies, or both, the evidence of interstudy

heterogeneity was for most endpoints greatly reduced (table 3, table 5), and the values of the  $I^2$  statistic were also lower. The causes of the heterogeneity are not known, although perhaps the multiplicity of lifestyle and medical risk factors in these studies could have a role. However, little evidence exists for the modifying effect of lifestyle and medical factors on excess relative risk associated with radiation. The different target tissues or their surrogates used in specific studies (eg, whole heart, coronary artery, left anterior descending artery, lung for ischaemic heart disease, carotid, thyroid gland, salivary gland, whole brain, Willis Circle arteries for cerebrovascular disease; supplement S3 tables S3.4, S3.5, S3.6) might have differences in radiosensitivity, and this might also contribute to the heterogeneity.

A limitation of our analysis, as with all meta-analysis, is that the effects of key aspects such as dose, dose rate, or age at exposure were only measured at the level of the study. We adjusted for these and other factors for each study via meta-regression (table 2, table 4). Inevitably such meta-regressions are quite generalised, amounting to a type of ecological analysis, and might not adequately control for the effects of these factors.

Many of the studies of medical exposure (supplement S3 table S3.4) had a substantial amount of information on standard lifestyle and medical risk factors for cardiovascular disease. However, information was more limited in the lower dose occupational or environmental studies (supplement table S3.5). In the lower dose studies of the Japanese atomic bomb survivors,<sup>20</sup> Mayak workers,<sup>64-66</sup> and a few other occupationally<sup>90-94</sup> and environmentally exposed<sup>95-96</sup> groups, substantial information was available on lifestyle factors (supplement S3 table S3.5). In most groups that were exposed to radiation, lifestyle risk factors had little or no evidence of interacting with cardiovascular disease risk related to radiation.<sup>13 19 20 54 56 57 59-63 67 90 91 93 95 97-99</sup> The types of chemotherapy that were likely to have been administered are more problematic because some (eg, anthracyclines) are known to be cardiotoxic. As discussed above, the data<sup>56-61</sup> do not suggest that these modify radiation risk. Nevertheless, such factors could confound the associations observed in relation to radiation exposure, although such confounding is unlikely. In some of these studies such information was collected and used in the analysis, and we give this information in supplement S3 tables S3.4-S3.6.

The heterogeneity in outcomes, and how they are defined and aggregated, is a potential problem in conducting reviews of this sort. We used both incidence and mortality data, in some cases within the same cohort. Our analysis highlights differences in meta excess relative risk per Gy of these factors (table 3). Mortality data could be more reliable because disease diagnosis (by a physician knowing the patient's history) could vary with dose. Although this issue is unlikely to affect the data relating to the Mayak workers,<sup>100</sup> this variability is of more concern in the Russian Chernobyl recovery workers.<sup>101-103</sup>

Conversely, if ascertainment of incidence can be done in a uniform way, incidence data are to be preferred because mortality data are intrinsically less reliable. Arguably the variety of different dose metrics used in each study is problematic (supplement S3 tables S4, S5). Wherever possible, we used the absorbed dose (which generally is unweighted dose) to the relevant organ (heart for ischaemic heart disease, dose to carotid artery, or salivary gland for cerebrovascular disease). However, only effective dose is given in some studies. Supplement S3 table S3.6 shows the risks in some therapeutic studies where alternative measures of dose were used. This variety does not suggest marked variation in risk by dose metric within a study and endpoint; however, this could contribute to interstudy heterogeneity.

We used an adaptation of the ROBINS-I system,<sup>37</sup> which we modified to make slightly more quantitative, in relation to the extent of likely bias. However, we regard this adapted ROBINS-I system as still quite subjective. For that reason, we also used the much more algorithmic (and less subjective) assessment of study quality, a slight modification of a system that was previously employed.<sup>21</sup> The advantage of these two scoring systems is that they yield semiquantitative scores of study quality. The United Nations Scientific Committee on the Effects of Atomic Radiation has outlined some general principles to be used in assessing study quality<sup>104</sup> but these are even less specific than the ROBINS-I system, and would not obviously result in a quantitative score.

### Strengths

Our systematic review and meta-analysis provides substantial advances on several earlier reviews, which also tended to concentrate on groups exposed to moderate and low doses of radiation.<sup>21 25 26</sup> Other more recent reviews have not been of systematic form.<sup>27 105</sup> We used state-of-the-art meta-analysis techniques to highlight possible contributions of dose level and dose fractionation. Another striking feature is that despite variation in quality of the individual studies (supplement S3 tables S3.1, S3.2), the overall inference was not much affected when attention was restricted to higher quality studies (table 4, table 5; supplement S3 fig S3.1, S3.2). Evidence has indicated selection bias, particularly for ischaemic heart disease, although if attention is restricted to lower dose (<0.5 Gy) or fractionated data, evidence of bias was much weaker, as also when using higher quality data only (table 3, table 5). The screening process was conducted independently by two authors (MPL, NH), so that omission of studies is unlikely. The data abstraction process was conducted independently by two authors (MPL, KA) and the results were also checked by a third (NH), so that errors in this process are even less likely.

### Conclusions

Our systematic review and meta-analysis supports an association between acute high dose and chronic low dose radiation exposure and most types of

cardiovascular disease. Low dose and low dose rate exposure tend to be associated with higher risk per unit dose. Although heterogeneity complicates a causal interpretation of these findings, this heterogeneity is markedly reduced if attention is restricted to higher quality studies or to studies at lower dose or dose rate. Our findings suggest that radiation detriment might have been significantly underestimated, implying that radiation protection and optimisation at low doses should be rethought. The possible mechanisms for risk at low doses and low dose rates are, in contrast to the situation at higher doses and dose rates, relatively poorly understood, thus underscoring a crucial need for further research in this area.<sup>106 107</sup> Further research is also needed to assess modifications of radiation effect by other lifestyle and medical risk factors.

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The lead author (MPL) affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained.

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**Web appendix 1:** Supplements S1-S4

**Web appendix 2:** Supplement S5 - Cardiovascular disease datafiles, quality coding, R script files